

técnica

Revista de Engenharia

Publicação da Associação dos Estudantes
do Instituto Superior Técnico

2 CARTA DOS EDITORES

3 EDITORIAL

artigos por convite

5 NORMAS DE CONSTRUÇÃO DAS CIDADES
— O SISTEMA ESTRATÉGICO — ESTOCÁSTICO
M.L. DA COSTA LOBO

15 O PROBLEMA DE WIENER-HOPF E ALGUMAS
DAS SUAS APLICAÇÕES
A. FERREIRA DOS SANTOS

artigos

29 FUNDAMENTOS DO VOO MAGNÉTICO.
A ANÁLISE DE UM MODELO ELEMENTAR
J.F. BORGES DA SILVA

35 A ESCOLHA DA CLASSE DE INTENSIDADE DOS PÁRA-RAIOS
PARA OS POSTOS DE TRANSFORMAÇÃO ALIMENTADOS POR
LINHAS AÉREAS A MÉDIA TENSÃO
DOMINGOS MOURA

47 APLICAÇÃO DA VOLTAMETRIA CÍCLICA À CARACTERIZAÇÃO
DE PROCESSOS DE ELÉCTRODO
M.A.G. MARTINS e C.A.C. SEQUEIRA

61 INVESTIGAÇÃO E DESENVOLVIMENTO.
A CONTRIBUIÇÃO DO CURSO DE MINAS DO I.S.T.
LUÍS AIRES-BARROS

67 MAPAS DE MECANISMOS COMPETITIVOS DE
RESPOSTA DE UM SISTEMA
M. AMARAL FORTES

notas científicas

73 A EQUAÇÃO RELATIVISTA DO MOVIMENTO
DE UMA BARRA NÃO RÍGIDA
A. BROTAS e J.C. FERNANDES

76 MODEL VALIDATION IN THE ESTIMATION
OF NON-STATIONARY SPATIAL PHENOMENA
HENRIQUE GARCIA PEREIRA

79 A VARIABLE METRIC PROJECTION METHOD FOR
MINIMIZING GENERAL FUNCTION SUBJECT
TO UPPER AND LOWER BOUNDS ON THE VARIABLES
MÁRIO J.A. LANÇA

82 THE ESTIMATION OF TRANSFER FUNCTION IN NATURAL
RESOURCES EVALUATION
FERNANDO DE OLIVEIRA MUGE

MODEL VALIDATION IN THE ESTIMATION OF NON-STATIONARY SPATIAL PHENOMENA

HENRIQUE GARCIA PEREIRA, * Prof. IST

ABSTRACT

Experimental results of the application of a covariance (or variogram) model validation method for non-stationary phenomena are presented for the case of estimation of S grades in a pyrite deposit.

RESUMO

Apresentam-se os resultados experimentais da aplicação, à estimação do teor em S numa jazida de pirites, de um método de validação de um modelo de covariância (ou variograma) para fenómenos não-estacionários.

INTRODUCTION

Many spatial phenomena occurring in Nature may be described by means of Random Functions. Characteristic variables of these phenomena (measured in some experimental points) are viewed as the realization of a Random Function whose covariance must be inferred to allow the BLUE (best linear unbiased estimation) of the variable in unknown points or surfaces. The inference of that covariance from the experimental data is not possible from one single realization of the phenomenon, when a stationarity hypothesis does not hold.

For a non-stationary phenomenon, the traditional approach to deal with estimation problems is to split the variable $Z(x)$ (where x are the coordinates of a point in one or two dimensions) in two terms:

$$Z(x) = m(x) + Y(x)$$

Where $m(x) = E[Z(x)]$ is the **drift**, an ordinary function that accounts for the regional component of the variable and $Y(x)$ is the **fluctuation**, a stationary random function with zero mean and stationary covariance that models the variability of the local component of the variable.

If the covariance $C(\bar{h}) = E[Y(x + \bar{h})Y(x)]$ was known, the problem of estimation of the variable in an arbitrary point x_0 could be solved on the grounds

of the Universal Kriging theory (MATHERON, 1970), using the system:

$$\begin{cases} \lambda^\beta C(x_\alpha - x_\beta) = C(x_\alpha - x_0) + \mu_e f^e(x_0) & \alpha = 1, 2, \dots, n \\ \lambda^\alpha f^e(x_\alpha) = f^e(x_0) & e = 0, 1, \dots, k \end{cases} \quad [1]$$

Where λ^α and λ^β are the weights to apply to the values of the variable in the n experimental points α in order to build up the estimator $\hat{Z}(x_0) = \lambda^\alpha Z(x_\alpha)$; μ_e are Lagrange parameters and f^e are monomials, depending on the degree k of the drift ($m(x) = a_e f^e(x)$). The Einstein summation convention is used in the sequel.

The kriging variance (minimum variance of $\hat{Z}(x_0) - Z(x_0)$) is calculated by:

$$\sigma_k^2 = C(0) + \mu_e f^e(x_0) - \lambda^\alpha C(x_\alpha - x_0) \quad [2]$$

MODEL VALIDATION

In order to apply system [1], a covariance model must be found. In Geostatistics (MATHERON, 1970), instead of covariance, it is worth to use σ variogram function, as many phenomena display an infinite dispersion. The variogram $\gamma(h)$ is calculated by:

$$\begin{aligned} \gamma(\bar{h}) &= \text{VAR} [Y(x + \bar{h}) - Y(x)] \\ &= E [Y(x + \bar{h}) - Y(x)]^2 \end{aligned}$$

and system [1] holds if $C(\bar{h})$ is substituted by $-\gamma(\bar{h})$.

But the variogram function to which one has experimental access is $\gamma_B(h) = \text{VAR} [Z(x + h) - Z(x)]$ (variogram of the original variable $Z(x)$) and, except in some particular cases (PEREIRA, 1983), the variogram of $Y(x)$ can not be calculated from data, because considerable biases are introduced.

So, an a priori model for the variogram function must be set, and parameters of the model must be calculated. In this note, a method of model validation

FIG. 1 — Experimental configuration of S grades (%)

is applied to experimental data of S grades in a pyrite deposit (CVRM, 1978).

The method consists in the computation of estimators $\hat{Z}(x_\alpha)$ (using $\hat{Z}(x_\alpha) = \lambda^\beta Z(x_\beta)$, $\forall \beta \neq \alpha$) in all experimental points where a value of $Z(x_\alpha)$ exists. Those estimators are obtained from system [1], using the a priori model for $\gamma(\bar{h})$. The parameters of the model are chosen in a practical range whose limits are suggested by a preliminary analysis of experimental data (variograms of $Z(x)$ and of $Z(x) - \hat{m}(x)$ being the basis of this analysis) and by all available qualitative information about the spatial development of the phenomenon.

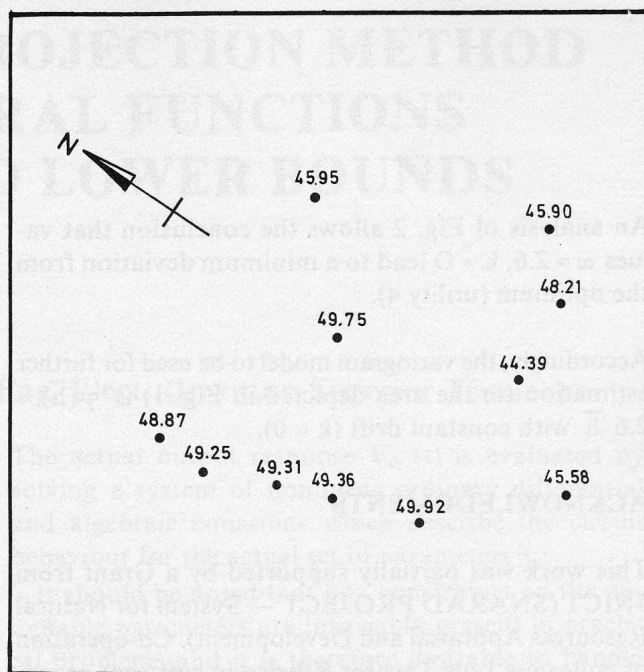
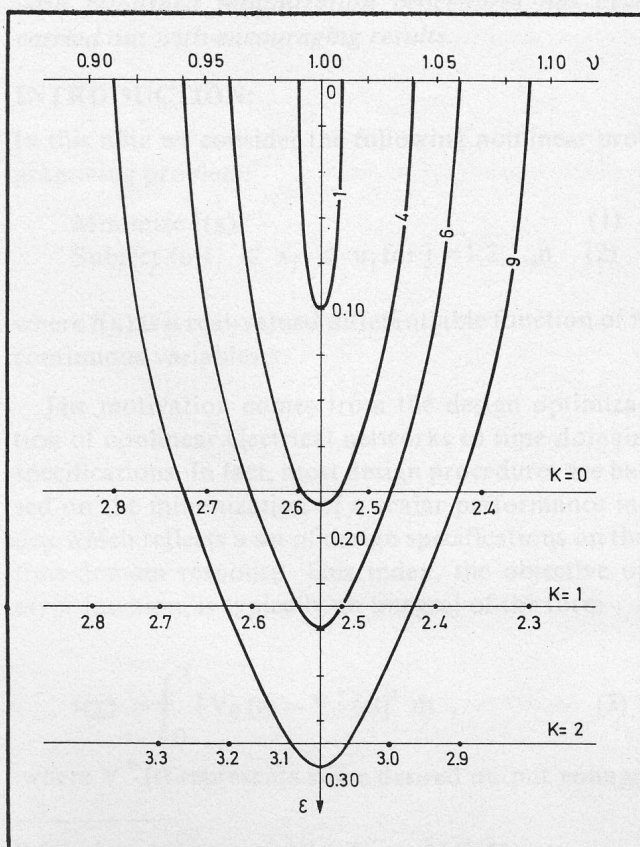
An utility surface is then built in function of two statistics:

- The Average error

$$\epsilon = \frac{1}{n} \sum_{\alpha} [\hat{Z}(x_\alpha) - Z(x_\alpha)]$$

- The Average Relative Quadratic error

$$\nu = \frac{1}{n} \sum_{\alpha} [\hat{Z}(x) - Z(x)]^2 / \sigma_k^2$$



The shape of indifference curves in the neighbourhood of the optimum (given by $\epsilon = 0, \nu = 1$) depends on the relative "preference" for each kind of error (the first one, ϵ , accounts for the average under or over estimation of the variable and the second one, ν , measures the goodness of fit of kriging variances calculated from the model).

The model chosen for further estimation procedures is the one for which values of statistics ϵ and ν are closest to the optimum, in the utility surface space.

EXPERIMENTAL RESULTS

The above described method was applied to an experimental configuration of 11 samples of pyrite ore assayed for S grades in a level of a pyrite deposit (Fig. 1).

The variogram model tested was $\gamma(\bar{h}) = \omega \bar{h}$, for ω in the range [2,3] and for $k = 0, 1, 2$ (degree of the drift).

In the utility surface shown in Fig. 2 (each indifference curve being labeled by a code 1,4,6,9 denoting a growing deviation from the optimum $\epsilon = 0, \nu = 1$), the experimental points for each pair of the model's parameters ω, k were plotted in function of ϵ and ν . It is worth noting that ϵ is constant for each k value (propriety of linear variogram estimation) and values of ω are figures appearing in each straight line $\epsilon = \text{constant}$.

FIG. 2 — Plot of experimental points (ω, k) in the utility surface space

An analysis of Fig. 2 allows the conclusion that values $\omega = 2.6$, $k = 0$ lead to a minimum deviation from the optimum (utility 4).

Accordingly, the variogram model to be used for further estimation (in the area depicted in Fig. 1) is $\gamma(\bar{h}) = 2.6 \bar{h}$ with constant drift ($k = 0$).

ACKNOWLEDGMENTS

This work was partially supported by a Grant from JNICT (SNARAD PROJECT — System for Natural Resources Appraisal and Development). Co-operation of colleague Luís Tavares Ribeiro in software prepa-

ration is appreciated. The author is indebted to Professor Quintino Rogado for his encouragement.

REFERENCES

- CVRM (1978) — Reavaliação Geoestatística e amostragem dos jazigos de MOINHO e FEITAIS, unpublished internal report, 144 p.
- MATHERON, G. (1970) — La théorie des variables regionalisées et ses applications. Les Cahiers du Centre de Morphologie Mathématique de Fontainebleau, Fascicule 5, 211 p.
- PEREIRA, H.G. (1983) — Estimação de funções aleatórias não estacionárias. in Encontros sobre Métodos Quantitativos Aplicados às Variáveis Regionalizadas, INIC, p. 115-131.

